

System Identification

Control Engineering EN, 3rd year B.Sc.
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Part II

Mathematical Background: Linear Regression. Probability Theory and Statistics

Motivation

Many methods for system identification require **linear regression** and some concepts from **probability theory and statistics**. We will discuss these mathematical tools here.

In this part some notation (e.g. x , A) has a different meaning than in the rest of the course.

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Motivating example

We study the **yearly income** (in EUR) of a person based on their **education level** and **job experience** (both measured in years).

We are given a dataset of tuples (education level, job experience, yearly income) from a representative set of persons. The goal is to **predict** the income of any other person, who is not in the dataset, by knowing how educated and experienced they are.

Regression problem: Elements

Problem elements:

- A collection of known samples $y(k) \in \mathbb{R}$, indexed by $k = 1, \dots, N$: y is the *regressed variable*.
- For each k , a known vector $\varphi(k) \in \mathbb{R}^n$: contains the *regressors* $\varphi_i(k)$, $i = 1, \dots, n$, $\varphi(k) = [\varphi_1(k), \varphi_2(k), \dots, \varphi_n(k)]^\top$.
- An **unknown** parameter vector $\theta \in \mathbb{R}^n$.

Linear model of the regressed variable:

$$\begin{aligned} y(k) &= \varphi^\top(k) \theta + e(k) \\ &= [\varphi_1(k), \varphi_2(k), \dots, \varphi_n(k)] \cdot [\theta_1, \theta_2, \dots, \theta_n]^\top + e(k) \\ &= \left[\sum_{i=1}^n \varphi_i(k) \theta_i \right] + e(k) \end{aligned}$$

where $e(k)$ is the model error

A remark on vector notation

All vector variables are column by default, following standard control notation. Often we write them as transposed rows, to save vertical space:

$$Q = \begin{bmatrix} q_1 \\ q_2 \\ q_3 \end{bmatrix} = [q_1, q_2, q_3]^\top$$

Note also: $Q^\top = [q_1, q_2, q_3]$.

For instance, in the linear model formula:

$$y(k) = [\varphi_1(k), \varphi_2(k), \dots, \varphi_n(k)] \cdot [\theta_1, \theta_2, \dots, \theta_n]^\top + e(k) = \varphi^\top(k)\theta + e(k)$$

$\varphi(k)$ is originally column, but must be transformed into a row to make the inner product work. So we transpose it: $\varphi^\top(k)$, obtaining the row $[\varphi_1(k), \dots]$ (note here there is no transpose). On the other hand, θ must be a column in the inner product, hence it is not transposed; however, for convenience we write it as a transposed row, i.e. $[\theta_1, \dots]^\top$ (note the transpose!).

Regression problem: Objective

Objective: identify the behavior of the regressed variable from the data.

In more detail: Find values for the parameters θ so that the approximated variable $\hat{y}(k) = \varphi^\top(k)\theta$ is as close as possible to the true $y(k)$ for all k ; equivalently, so that model errors $e(k)$ are as small as possible.

We will clarify what "as close/small as possible" means later.

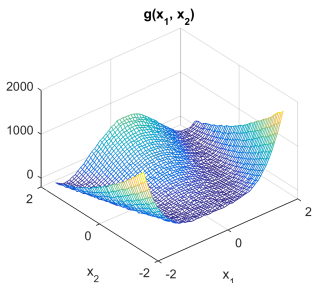
Linear regression is classical and very common, e.g. it was used by Gauss to compute the orbits of the planets in 1809.

Regression problem: Two major uses

- 1 k is a time variable, and we wish to model the time series $y(k)$.
- 2 k is just a data index, and $\varphi(k) = \phi(x(k))$ where x is an input of some unknown function g . Then $y(k)$ is the corresponding output (possibly corrupted by noise), and the goal is to identify a model of g from the data.

This problem is called *function fitting*, *function approximation*, or *supervised learning*.

unknown $g(x)$



$\hat{g}(x)$



Function approximation: Basis functions

For function approximation, the regressors $\phi_i(x)$ in:

$$\phi(x(k)) = [\phi_1(x(k)), \phi_2(x(k)), \dots, \phi_n(x(k))]^\top$$

are also called *basis functions*.

Function approximation: Formalizing the motivating example

We study the **yearly income** y (in EUR) of a person based on their **education level** x_1 and **job experience** x_2 (both measured in years).

We are given a set of tuples $(x_1(k), x_2(k), y(k))$ from a representative set of persons. The goal is to **predict** the income of any other person by knowing how educated (x_1) and experienced (x_2) they are.

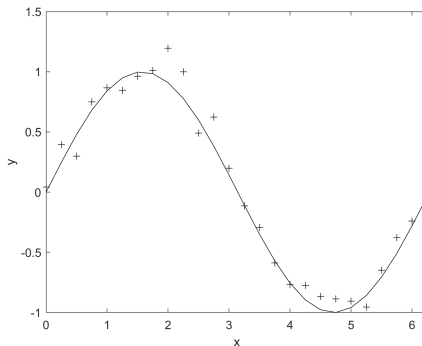
- Take basis functions $\phi(x) = [x_1, x_2, 1]^T$. So we expect the income to behave like $\theta_1 x_1 + \theta_2 x_2 + \theta_3 = \phi^T(x)\theta$, growing linearly with education and experience (from some minimum level). Regression involves finding the parameters θ in order to **best fit** the given data.
- Reality is of course more complicated... so we would likely need more input variables, better basis functions, etc.

Function approximation: Motivating example 2

We study the **reaction time** y (in ms) of a driver based on their **age** x_1 (in years) and **fatigue** x_2 (e.g. on a scale from 0 to 1).

We are given a set of tuples $(x_1(k), x_2(k), y(k))$ from a representative set of persons of various ages and stages of fatigue. The goal is to **predict** the reaction time of any other person by knowing how old (x_1) and tired (x_2) they are.

Function approximation: Running example



Approximate $\sin(x)$ over interval $[0, 2\pi]$. Note: No access to underlying function, only noisy samples! We'll use this as a running example to illustrate the concepts.

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Regressors example 1: Polynomial of k

Suitable for time series modeling.

$$\begin{aligned} y(k) &= \theta_1 + \theta_2 k + \theta_3 k^2 + \dots + \theta_n k^{n-1} \\ &= \begin{bmatrix} 1 & k & k^2 & \dots & k^{n-1} \end{bmatrix} \begin{bmatrix} \theta_1 \\ \theta_2 \\ \theta_3 \\ \dots \\ \theta_n \end{bmatrix} \\ &= \varphi^\top(k) \theta \end{aligned}$$

Regressors example 2: Polynomial of x

Suitable for function approximation. For instance,
polynomial of degree 2 with two input variables $x = [x_1, x_2]^T$:

$$y(k) = \theta_1 + \theta_2 x_1(k) + \theta_3 x_2(k) + \theta_4 x_1^2(k) + \theta_5 x_2^2(k) + \theta_6 x_1(k)x_2(k)$$

$$= \begin{bmatrix} 1 & x_1(k) & x_2(k) & x_1^2(k) & x_2^2(k) & x_1(k)x_2(k) \end{bmatrix} \begin{bmatrix} \theta_1 \\ \theta_2 \\ \theta_3 \\ \theta_4 \\ \theta_5 \\ \theta_6 \end{bmatrix}$$

$$= \phi^T(x(k))\theta = \varphi^T(k)\theta$$



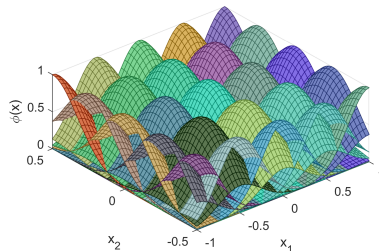
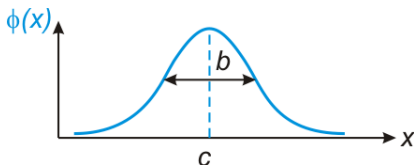
Connection: Project part 1

Regressors example 3: Gaussian basis functions

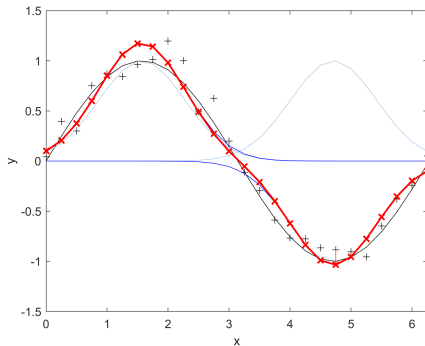
Suitable for function approximation:

$$\phi_i(x) = \exp \left[-\frac{(x - c_i)^2}{b_i^2} \right] \quad (1\text{-dim});$$

$$= \exp \left[-\sum_{j=1}^d \frac{(x_j - c_{ij})^2}{b_{ij}^2} \right] \quad (d\text{-dim})$$



Regressors: Running example



Since the data has a hill and a valley, **two RBFs** should suffice (see solution preview).

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Recall: regression objective

Objective: identify the behavior of the regressed variable from the data.

i.e., find values for the parameters θ so that the approximated variable $\hat{y}(k) = \varphi^\top(k)\theta$ is as close as possible to the true $y(k)$ for all k ; equivalently, so that model errors $e(k)$ are as small as possible.

Linear system

Writing the model for each of the N data points, we get a linear system of equations:

$$y(1) = \varphi_1(1)\theta_1 + \varphi_2(1)\theta_2 + \dots \varphi_n(1)\theta_n$$

$$y(2) = \varphi_1(2)\theta_1 + \varphi_2(2)\theta_2 + \dots \varphi_n(2)\theta_n$$

...

$$y(N) = \varphi_1(N)\theta_1 + \varphi_2(N)\theta_2 + \dots \varphi_n(N)\theta_n$$

Recall that in function approximation, $\varphi_i(k) = \phi_i(x(k))$

This system can be written in a *matrix form*:

$$\begin{bmatrix} y(1) \\ y(2) \\ \vdots \\ y(N) \end{bmatrix} = \begin{bmatrix} \varphi_1(1) & \varphi_2(1) & \dots & \varphi_n(1) \\ \varphi_1(2) & \varphi_2(2) & \dots & \varphi_n(2) \\ \dots & \dots & \dots & \dots \\ \varphi_1(N) & \varphi_2(N) & \dots & \varphi_n(N) \end{bmatrix} \cdot \begin{bmatrix} \theta_1 \\ \theta_2 \\ \theta_3 \\ \dots \\ \theta_n \end{bmatrix}$$

$$Y = \Phi\theta$$

with newly introduced variables $Y \in \mathbb{R}^N$ and $\Phi \in \mathbb{R}^{N \times n}$.

Least-squares problem

If $N = n$, the system can be solved with equality.

In practice, it is a good idea to use $N > n$, due e.g. to noise. In this case, the system can no longer be solved with equality, but only in an approximate sense.

- Error at k : $\varepsilon(k) = y(k) - \varphi^\top(k)\theta$,
error vector $\varepsilon = [\varepsilon(1), \varepsilon(2), \dots, \varepsilon(N)]^\top$.
- **Objective function** to be minimized:

$$V(\theta) = \frac{1}{2} \sum_{k=1}^N \varepsilon(k)^2 = \frac{1}{2} \varepsilon^\top \varepsilon$$

Least-squares problem

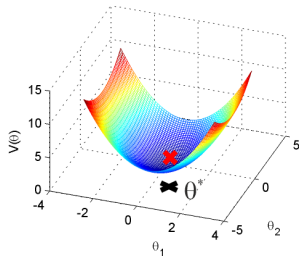
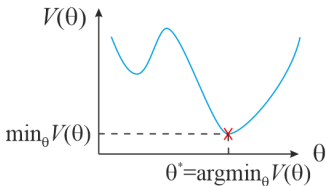
Find the parameter vector $\hat{\theta}$ that minimizes the objective function:

$$\hat{\theta} = \arg \min_{\theta} V(\theta)$$

Parenthesis: Optimization problem

Given a function V of variables θ , which may be the least-squares objective, or any other function:

find the *optimal function value* $\min_{\theta} V(\theta)$ and variable values $\theta^* = \arg \min_{\theta} V(\theta)$ that achieve the minimum.



Note that in the case of linear regression, we use the notation $\hat{\theta}$; while $\hat{\theta}$ is still the true solution to the optimization problem given the data, it is still an estimate because the data is noisy

Mathematical regression solution

For an easy derivation, start with:

$$Y = \Phi\theta$$

and left-multiply by $\Phi^\top$:

$$\Phi^\top Y = \Phi^\top \Phi\theta$$

Another left-multiplication by the inverse of $\Phi^\top \Phi$ leads to:

$$\hat{\theta} = (\Phi^\top \Phi)^{-1} \Phi^\top Y$$

Remarks:

- A better way of finding $\hat{\theta}$ is to write the optimization problem of minimizing the MSE, and solving it via setting the (matrix) derivative to 0.
- Matrix $\Phi^\top \Phi$ must be invertible. This boils down to a good choice of the model (order n , regressors φ) and having informative data.
- The optimal objective value is $V(\hat{\theta}) = \frac{1}{2}[Y^\top Y - Y^\top \Phi(\Phi^\top \Phi)^{-1} \Phi^\top Y]$.
- $(\Phi^\top \Phi)^{-1} \Phi^\top$ is the pseudo-inverse of Φ .

Alternative expression

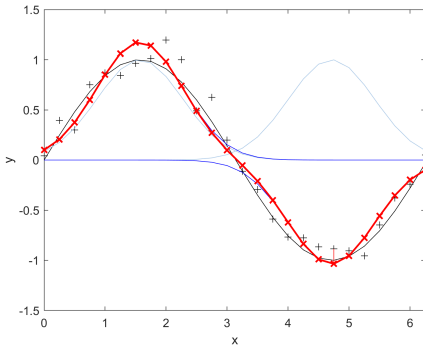
$$\Phi^T \Phi = \sum_{k=1}^N \varphi(k) \varphi^T(k), \Phi^T Y = \sum_{k=1}^N \varphi(k) y(k)$$

So the solution can be written:

$$\hat{\theta} = \left[\sum_{k=1}^N \varphi(k) \varphi^T(k) \right]^{-1} \left[\sum_{k=1}^N \varphi(k) y(k) \right]$$

Advantage: matrix Φ of size $N \times n$ no longer has to be computed, only smaller matrices and vectors are required, of size $n \times n$ and n , respectively.

Least-squares solution: Running example



Parameters θ_1, θ_2 **scale & flip the RBFs** appropriately, so that the **approximate solution** minimizes the sum of square distances between the real and approximated data.

Solving the linear system in practice

In practice both these inversion-based techniques perform poorly from a numerical point of view. Better algorithms exist, such as so-called orthogonal triangularization.

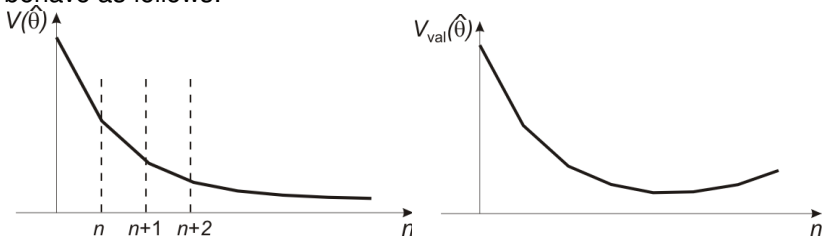
In most cases, **MATLAB** is competent in choosing a good algorithm. If Φ is stored in variable `PHI` and Y in `Y`, then the command to solve the linear system using matrix left division (backslash) is:

```
theta = PHI \ Y;
```

Better control of the algorithm is obtained by using function `linsolve` instead of matrix left division.

Model choice

Assume that given a model size n , we have a way to generate regressors $\varphi(k)$ that make the model more expressive (e.g., basis functions on a finer grid). Then we expect the objective function to behave as follows:



Remark: With noisy data, increasing n too much can lead to **overfitting**: good performance on the training data, but poor performance on other data. **Validation on a separate dataset** is essential in practice!

So we can grow n incrementally and stop when error V_{val} on the validation data starts growing.

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Analytical example: Estimating a scalar

Model:

$$y(k) = b = 1 \cdot b = \varphi(k)\theta$$

where $\varphi(k) = 1 \forall k, \theta = b$.

For all N data points:

$$y(1) = \varphi(1)\theta = 1 \cdot b$$

...

$$y(N) = \varphi(N)\theta = 1 \cdot b$$

In matrix form:

$$\begin{bmatrix} y(1) \\ \vdots \\ y(N) \end{bmatrix} = \begin{bmatrix} 1 \\ \vdots \\ 1 \end{bmatrix} \theta$$

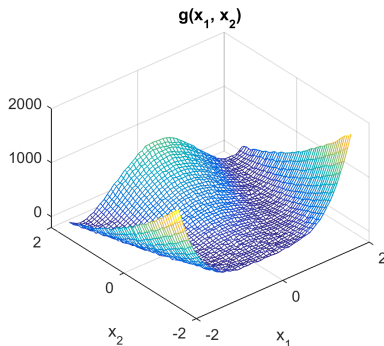
$$Y = \Phi\theta$$

Analytical example: Estimating a scalar (continued)

$$\begin{aligned}\hat{\theta} &= (\Phi^T \Phi)^{-1} \Phi^T Y \\ &= \left(\begin{bmatrix} 1 & \dots & 1 \end{bmatrix} \begin{bmatrix} 1 \\ \vdots \\ 1 \end{bmatrix} \right)^{-1} \begin{bmatrix} 1 & \dots & 1 \end{bmatrix} \begin{bmatrix} y(1) \\ \vdots \\ y(N) \end{bmatrix} \\ &= N^{-1} \begin{bmatrix} 1 & \dots & 1 \end{bmatrix} \begin{bmatrix} y(1) \\ \vdots \\ y(N) \end{bmatrix} \\ &= \frac{1}{N} (y(1) + \dots + y(N))\end{aligned}$$

Intuition: Estimate is the average of all measurements, filtering out the noise.

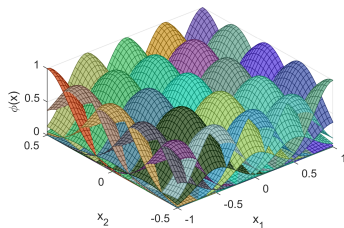
Example: Approximating the “banana” function



- Function $g(x_1, x_2) = (1 - x_1)^2 + 100[(x_2 + 1.5) - x_1^2]^2$, called Rosenbrock's banana function (unknown to the algorithm).
- **Approximation data:** 200 input points (x_1, x_2) , randomly distributed over the space $[-2, 2] \times [-2, 2]$; and corresponding outputs $y = g(x_1, x_2)$, **affected by noise**.
- **Validation data:** a uniform grid of 31×31 points over $[-2, 2] \times [-2, 2]$ with their corresponding (noisy) outputs.

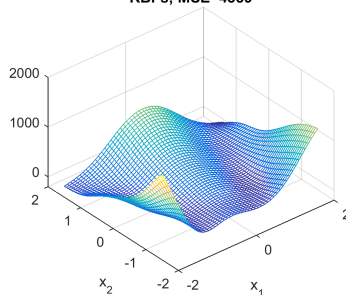
Banana function: Results with radial basis functions

Recall radial basis functions:



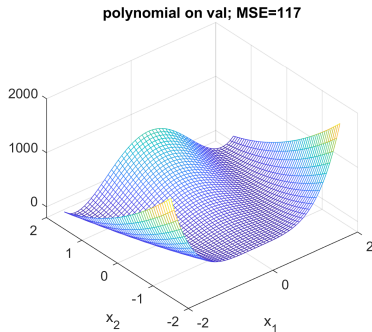
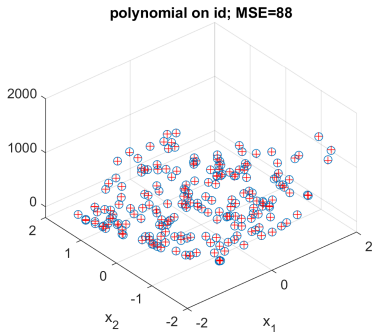
Results with 6×6 RBFs, with centers on an equidistant grid and width equal to the distance between two centers:

RBFs; MSE=4359



Banana function: Results with polynomial

Polynomial of degree 4 in the two input variables (15 parameters):

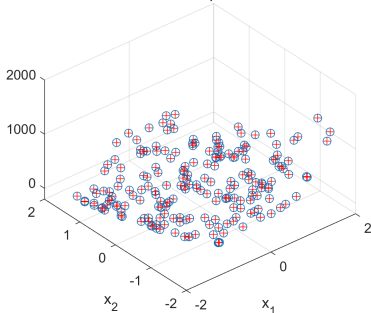


Banana function: Overfitting example

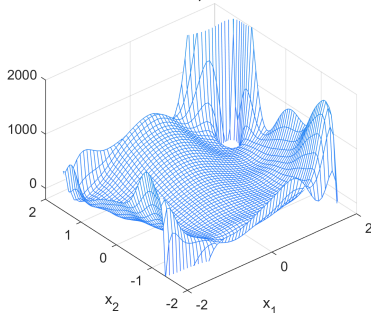
Again polynomial, but this time of degree 13.

The MSE on the identification data is smaller than for degree 4 above. However, on the validation data:

overfit on id; MSE=39



overfit on val; MSE=625363



Very bad result! Overly large degree leads to overfitting.

Summary linear regression

- Linear models for function and time series approximation.
- Polynomial and radial basis functions.
- Writing the problem as a linear system of equations, mathematical objective, and solution (theoretical and practical).
- Model choice and overfitting.
- Examples: estimating a scalar and the banana function.

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Discrete random variable

Consider a set $\mathcal{X}$ containing possible values x . A corresponding random variable X is described by the probability mass function:

Definition

The **probability mass function** (PMF) of X is the list of the probabilities of all individual values $p(x_0), p(x_1), \dots$. The probabilities must be positive, $p(x) \geq 0 \forall x$, and must sum up to 1:

$$P(X \in \mathcal{X}) = \sum_{x \in \mathcal{X}} p(x) = 1.$$

Example: The number rolled with a dice is a discrete random variable, with six possible values $\mathcal{X} = \{1, 2, \dots, 6\}$. For a fair dice, the PMF is $p(x) = 1/6$ for all x , from 1 to 6.

Note that $\mathcal{X}$ may contain finitely or infinitely many elements, but in the latter case, the set must be countable (Intuition: elements can be listed / indexed by the natural numbers $0, 1, 2, \dots$).

Motivating example for continuous random variables

We wish to characterize the angle at which a roulette wheel ends up, as a fraction of a full turn – a continuous variable: $x \in [0, 1]$. In this case, $\mathcal{X} = [0, 1]$. For a fair roulette, each value x has equal probability.

Let us try to define a PMF; it should have the same value everywhere. Denote this value by c ; since there are infinitely many values in any finite-length interval in $[0, 1]$, any nonzero value of c would lead to infinitely large probability for any such interval! So the only value of $c = p(x)$ that can work is 0, but this has no useful meaning – therefore, we cannot use a PMF to describe this case.

Meaningful probabilities can only be defined for intervals.

Continuous random variable

Consider an interval $\mathcal{X} \subseteq \mathbb{R}$ of real numbers, with each individual real value denoted x . A corresponding, continuous-valued random variable X is described by the probability density function:

Definition (semi-formal)

The **probability density function** (PDF) of X is a function $f : \mathcal{X} \rightarrow [0, \infty)$, such that the probability of obtaining a value in some subinterval $[a, b] \subseteq \mathcal{X}$ is:

$$P(X \in [a, b]) = \int_a^b f(x) dx$$

The PDF must satisfy $P(X \in \mathcal{X}) = \int_{\mathcal{X}} f(x) dx = 1$.

Remark: $\mathcal{X}$ might also be the full set of real numbers.

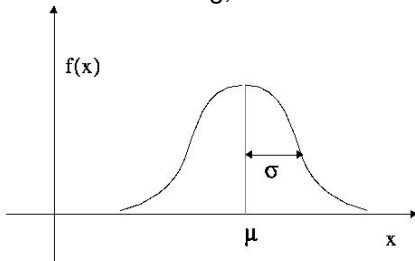
Example 1: Uniform PDF for roulette wheel

Let us return to the roulette wheel. Angular fraction of a full turn:
 $x \in [0, 1]$, and $\mathcal{X} = [0, 1]$.

To have uniform probabilities so as to represent a fair roulette, we want $P(x \in [a, b]) = b - a$, which means PDF $f(x) = 1$.

Example 2: Gaussian PDF

Similar in shape, but not in meaning, with Gaussian basis functions.



$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right)$$

Parameters: mean μ , and variance σ^2 (their meaning is clarified later)

The Gaussian distribution arises very often in nature: e.g., distribution of IQs in a human population. It is also called the normal distribution, and denoted $\mathcal{N}(\mu, \sigma^2)$.

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Expected value

Definition

$$E\{X\} = \mu_X = \begin{cases} \sum_{x \in \mathcal{X}} p(x)x & \text{for discrete random variables} \\ \int_{\mathcal{X}} f(x)x & \text{for continuous random variables} \end{cases}$$

Intuition: the average of all values, weighted by their probability; the value “expected” beforehand given the probability distribution.

The expected value is also called *mean*, or *expectation*.

Examples:

- For a fair dice with X the value of a face,
 $E\{X\} = \frac{1}{6}1 + \frac{1}{6}2 + \dots + \frac{1}{6}6 = 7/2$.
- If X has a Gaussian PDF $\mathcal{N}(\mu, \sigma^2)$, then $E\{X\} = \mu$. Note constant μ vs. mean of X , μ_X .

Expected value of a function

Consider a function $g : \mathcal{X} \rightarrow \mathbb{R}$, depending on some random variable X . Then $g(X)$ is itself a random variable, and:

$$\mathbb{E} \{g(X)\} = \begin{cases} \sum_{x \in \mathcal{X}} p(x)g(x) & \text{if discrete} \\ \int_{x \in \mathcal{X}} f(x)g(x) & \text{if continuous} \end{cases}$$

Example: Say we play a game of dice where face 6 gains 10 dollars, and other faces gain nothing. We have $g(6) = 10$ and $g(x) = 0$ for all other x . The expected value of this game is $\frac{1}{6}0 + \dots + \frac{1}{6}0 + \frac{1}{6}10 = 10/6$ dollars.

Variance

Definition

$$\text{Var} \{X\} = \sigma_X^2 = \boxed{\text{E} \{(X - \mu_X)^2\}} = \text{E} \{X^2\} - \mu_X^2$$

Intuition: the “spread” of the random values around the expectation.

$$\begin{aligned} \text{Var} \{X\} &= \begin{cases} \sum_{x \in \mathcal{X}} p(x)(x - \mu_X)^2 & \text{if discrete} \\ \int_{x \in \mathcal{X}} f(x)(x - \mu_X)^2 & \text{if continuous} \end{cases} \\ &= \begin{cases} \sum_{x \in \mathcal{X}} p(x)(x^2 - \mu_X^2) & \text{if discrete} \\ \int_{x \in \mathcal{X}} f(x)(x^2 - \mu_X^2) & \text{if continuous} \end{cases} \end{aligned}$$

Quantity $\sigma_X = \sqrt{\text{Var} \{X\}}$ is called *standard deviation*.

Examples:

- For a fair dice, $\text{Var} \{X\} = \frac{1}{6}1^2 + \frac{1}{6}2^2 + \dots + \frac{1}{6}6^2 - (7/2)^2 = 35/12$.
- If X has a Gaussian PDF $\mathcal{N}(\mu, \sigma^2)$, then $\text{Var} \{X\} = \sigma^2$.

Probability: Independence

Definition

Two random variables X and Y are called **independent** if:

- in the continuous case, $f_{X,Y}(x, y) = f_X(x) \cdot f_Y(y)$.
- in the discrete case, $p_{X,Y}(x, y) = p_X(x) \cdot p_Y(y)$.

where $f_{X,Y}$ denotes the joint PDF of the vector (X, Y) , f_X and f_Y are the PDFs of X and Y , and similarly for the PMFs p .

Examples:

- The event of rolling a 6 with a dice is independent of the event of getting a 6 at the previous roll (or, indeed, any other value and any previous roll).
- The event of rolling *two consecutive* 6-s, however, is not independent of the previous roll!

(Incidentally, failing to understand the first example leads to the so-called gambler's fallacy. Having just had a long sequence of bad—or good—games at the casino, means nothing for the next game!)

Covariance

$$\text{Var} \{X\} = E \{(X - \mu_X)^2\} = E \{(X - \mu_X)(X - \mu_X)\}$$

Replace X by Y in the second term.

Definition

$$\text{Cov} \{X, Y\} = E \{(X - \mu_X)(Y - \mu_Y)\} = E \{(X - E \{X\})(Y - E \{Y\})\}$$

Intuition: how much the two variables “change together” (positive if they change in the same way, negative if they change in opposite ways).

Remark: $\text{Var} \{X\} = \text{Cov} \{X, X\}$.

Uncorrelated variables

Definition

Random variables X and Y are **uncorrelated** if $\text{Cov} \{X, Y\} = 0$. Otherwise, they are **correlated**.

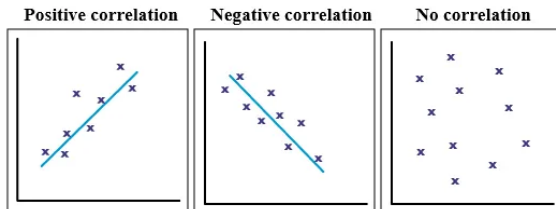
Examples:

- The education level of a person is correlated with their income.
- Hair color is uncorrelated with income (or at least it should be, ideally).

Remarks:

- If X and Y are independent, they are uncorrelated.
- But the reverse is not necessarily true! Variables can be uncorrelated and still dependent.

Covariance intuition



X and Y axes of each plot are the two random variables which we study.

- Positively correlated (covariance > 0): when X increases, Y increases.
- Negatively correlated (covariance < 0): when X increases, Y decreases.
- Uncorrelated (covariance $= 0$): no relation.

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Vectors of random variables

Consider a vector $\mathbf{X} = [X_1, \dots, X_N]^\top$ where each X_i is a continuous, real random variable. This vector has a *joint PDF* $f(\mathbf{x})$, with $\mathbf{x} \in \mathbb{R}^N$.

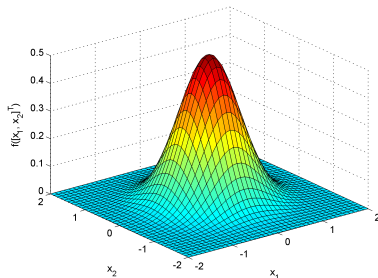
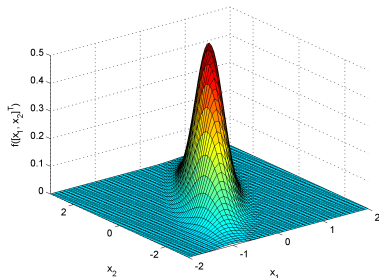
Definitions

Expected value and covariance matrix of $\mathbf{X}$:

$$\begin{aligned} \mathbb{E} \{\mathbf{X}\} &= \boldsymbol{\mu}_{\mathbf{X}} := [\mathbb{E} \{X_1\}, \dots, \mathbb{E} \{X_N\}]^\top = [\mu_{X_1}, \dots, \mu_{X_N}]^\top \in \mathbb{R}^N \\ \text{Cov} \{\mathbf{X}\} &= \Sigma_{\mathbf{X}} := \begin{bmatrix} \text{Cov} \{X_1, X_1\} & \text{Cov} \{X_1, X_2\} & \cdots & \text{Cov} \{X_1, X_N\} \\ \text{Cov} \{X_2, X_1\} & \text{Cov} \{X_2, X_2\} & \cdots & \text{Cov} \{X_2, X_N\} \\ \cdots & \cdots & \cdots & \cdots \\ \text{Cov} \{X_N, X_1\} & \text{Cov} \{X_N, X_2\} & \cdots & \text{Cov} \{X_N, X_N\} \end{bmatrix} \\ &= \mathbb{E} \left\{ (\mathbf{X} - \boldsymbol{\mu}_{\mathbf{X}})(\mathbf{X} - \boldsymbol{\mu}_{\mathbf{X}})^\top \right\} \in \mathbb{R}^{N,N} \end{aligned}$$

Remarks: $\text{Cov} \{X_i, X_i\} = \text{Var} \{X_i\}$. Also, $\text{Cov} \{X_i, X_j\} = \text{Cov} \{X_j, X_i\}$, so matrix $\Sigma_{\mathbf{X}}$ is symmetrical.

Example: Multivariate Gaussian (normal)



The PDF of a vector $\mathbf{X}$ with a Gaussian joint distribution can be written:

$$f(\mathbf{x}) = \frac{1}{(2\pi)^N \sqrt{\det(\Sigma)}} \exp \left(-(\mathbf{x} - \boldsymbol{\mu}) \Sigma^{-1} (\mathbf{x} - \boldsymbol{\mu})^\top \right)$$

parameterized by the vector mean $\boldsymbol{\mu}$ and covariance matrix Σ (assumed positive definite, so $\det(\Sigma) > 0$ and Σ^{-1} exists).

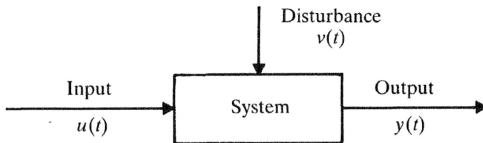
Stochastic process

Definition

A **stochastic process** $\mathbf{X}$ is a sequence of random variables $\mathbf{X} = (X_1, \dots, X_k, \dots, X_N)$.

It is in fact just a vector of random variables, with the additional structure that the index in the vector has the meaning of time step k .

In system identification, signals (e.g., inputs, outputs) will often be stochastic processes evolving over discrete time steps k .



Zero-mean white noise

Definition

The stochastic process $\mathbf{X}$ is **zero-mean white noise** if: $\forall k, E \{X_k\} = 0$ (zero-mean), and $\forall k, k' \neq k, \text{Cov} \{X_k, X_{k'}\} = 0$ (the values at different steps are uncorrelated). In addition, the variance $\text{Var} \{X_k\}$ must be finite $\forall k$.

Stated concisely using vector notation: mean $\mu_{\mathbf{X}} = E \{\mathbf{X}\} = \mathbf{0} \in \mathbb{R}^N$ and covariance matrix $\Sigma_{\mathbf{X}} = \text{Cov} \{\mathbf{X}\}$ is diagonal (with positive finite elements on the diagonal).

In system identification, noise processes often affect signal measurements, and we will sometimes assume that the noise is white and zero-mean.

Stationary process

Signal values at different time steps can be correlated (e.g. when they depend on the output of some dynamic system). Nevertheless, signals are sometimes required to be stationary, in the sense:

Definition

The stochastic process $\mathbf{X}$ is **stationary** if $\forall k, E\{X_k\} = \mu$, and $\forall k, k', \tau, \text{Cov}\{X_k, X_{k+\tau}\} = \text{Cov}\{X_{k'}, X_{k'+\tau}\}$.

The mean is the same at every step, whereas the covariance depends on only the relative positions of time steps, not their absolute positions.

Summary probability theory

- Discrete-valued random variables (RVs) and probability mass functions.
- Continuous-valued RVs and probability density functions.
- Expected value and variance of a RV.
- Independence, covariance, and correlation of two RVs.
- Vector of RVs, with expected value and covariance matrix.
- Stochastic process, zero-mean white noise, and stationary process.